# Variational Inference

CS772A: Probabilistic Machine Learning Piyush Rai

# Variational Inference (VI)

- Assume a latent variable model with data  ${m {\cal D}}$  and latent variables  ${m Z}$
- A simple setting might look something like this



This setting is just one example. VI is applicable in more general and more complex probabilistic models with and without latent variables

- Assume the likelihood is  $p(\mathcal{D}|Z, \Theta)$  and prior is  $p(Z|\Theta)$ . Want posterior over Z
- $\Theta = (\theta, \phi)$  denotes the other parameters that define the likelihood and the prior
- For now, assume  $\Theta$  is known and only Z is unknown (the  $\Theta$  unknown case later)
- Assume CP  $p(\mathbf{Z}|\mathbf{D}, \Theta)$  is intractable



# Variational Inference (VI)

• Assuming  $p(Z|\mathcal{D},\Theta)$  is intractable, VI approximates it by a distr  $q(Z|\phi)$  or  $q_{\phi}(Z)$ 



# Variational Inference (VI)

The optimization problem

$$\begin{split} \phi^* &= \operatorname{argmin}_{\phi} \operatorname{KL}[q_{\phi}(Z)||p(Z|\mathcal{D}, \Theta)] \\ &= \operatorname{argmin}_{\phi} \mathbb{E}_{q_{\phi}(Z)} \left[ \log q_{\phi}(Z) - \log \frac{p(\mathcal{D}|Z, \Theta)p(Z|\Theta)}{p(\mathcal{D}|\Theta)} \right] \\ &= \operatorname{argmin}_{\phi} \mathbb{E}_{q_{\phi}(Z)} [\log q_{\phi}(Z) - \log p(\mathcal{D}|Z, \Theta) - \log p(Z|\Theta)] + \log p(\mathcal{D}|\Theta) \\ \bullet \text{ Since } \log p(\mathcal{D}|\Theta) \text{ is independent of } \phi, \text{ the optimization problem becomes} \\ \phi^* &= \operatorname{argmin}_{\phi} \mathbb{E}_{q_{\phi}(Z)} [\log q_{\phi}(Z) - \log p(\mathcal{D}|Z, \Theta) - \log p(Z|\Theta)] \\ \phi^* &= \operatorname{argmin}_{\phi} \mathbb{E}_{q_{\phi}(Z)} [\log q_{\phi}(Z) - \log p(\mathcal{D}, Z|\Theta)] \\ \phi^* &= \operatorname{argmax}_{\phi} \mathbb{E}_{q_{\phi}(Z)} [\log p(\mathcal{D}, Z|\Theta) - \log q_{\phi}(Z)] = \operatorname{argmax} \mathcal{L}(\phi, \Theta) \\ \bullet \text{ Note that } \mathcal{L}(\phi, \Theta) \leq \log p(\mathcal{D}|\Theta) \text{ and is called "Evidence Lower Bound" (ELBO)} \end{split}$$

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## The ELBO

# • The ELBO is defined as $\mathcal{L}(\phi, \Theta) = \mathbb{E}_{q_{\phi}(Z)} [\log p(\mathcal{D}, Z | \Theta) - \log q_{\phi}(Z)]$ $= \mathbb{E}_{q_{\phi}(Z)} [\log p(\mathcal{D}, Z | \Theta)] + H[q_{\phi}(Z)]$

- Thus maximizing the ELBO w.r.t.  $\phi$  gives us a  $q_{\phi}(Z)$  which
  - Maximizes the expected joint probability of data and latent variables
  - Has a high entropy
- We can also write the ELBO as follows

#### $\mathcal{L}(\phi, \Theta) = \mathbb{E}_{q_{\phi}(Z)}[\log p(\mathcal{D}|Z, \Theta)] - \mathrm{KL}[q_{\phi}(Z)||p(Z|\Theta)]$

- Thus maximizing the ELBO w.r.t.  $\phi$  will give us a  $q_{\phi}(Z)$  which
  - Explains the data  $\mathcal{D}$  well, i.e., gives it large <u>expected</u> probability  $\mathbb{E}_q[\log p(\mathcal{D}|Z, \Theta)]$
  - Is close to the prior p(Z), i.e. is simple/regularized (small  $\mathrm{KL}[q_{\phi}(Z)||p(Z|\Theta))$ )



# Maximizing the ELBO

• We need to maximize the ELBO w.r.t.  $\phi$  (for now, assuming  $\Theta$  is known)

#### $\mathcal{L}(\phi, \Theta) = \mathbb{E}_{q_{\phi}(\mathbf{Z})}[\log p(\mathbf{\mathcal{D}}|\mathbf{Z}, \Theta)] - \mathrm{KL}[q_{\phi}(\mathbf{Z})||p(\mathbf{Z}|\Theta)]$

- The general approach to maximize ELBO is based on gradient-based methods
  - Assume some suitable/convenient form for  $q_{\phi}(Z)$ , e.g.,  $\mathcal{N}(Z|\mu, \Sigma)$  so  $\phi = (\mu, \Sigma)$
  - Maximize the ELBO w.r.t.  $\phi$  using gradient ascent

 $\phi_{t+1} = \phi_t + \eta_t \, \nabla_{\phi_t} \mathcal{L}(\phi, \Theta)$ 

 $\blacksquare$  Note: Expectations in ELBO and ELBO's gradients w.r.t.  $\phi$  may not be easy

- Will see methods to handle such issues later
- Assuming simple forms for  $q_{\phi}(Z)$  also helps (we can use random variable transformation methods to transform the simple form to more expressive ones will see later)

Unknown  $\Theta$  case later

# A Simple Illustration for VI

Assume a simple likelihood model

$$p(\boldsymbol{\mathcal{D}}|\boldsymbol{z}) = \prod_{n=1}^{N} \mathcal{N}(\boldsymbol{x}_{n}|\boldsymbol{z},\boldsymbol{\Sigma}) \propto \mathcal{N}(\overline{\boldsymbol{x}}|\boldsymbol{z},\frac{1}{N}\boldsymbol{\Sigma})$$

- $\hfill\blacksquare$  Suppose we want to estimate the posterior of the mean z
- Assuming a Gaussian prior on z and assuming  $\Sigma$  is known, the posterior can be computed analytically (because of conjugacy)
- Let's still try VI to see how well it does
- Figure shows VI result for three Gaussian forms for q(z)
  - Low-rank:  $q(z) = \mathcal{N}(z|\mu_z, \Sigma_z)$  where  $\Sigma_z = LL^{\mathsf{T}}$
  - Full-rank:  $q(z) = \mathcal{N}(z|\mu_z, \Sigma_z)$  with no constraint on  $\Sigma_z$
  - Mean-field:  $q(z) = q(z_1)q(z_2) = \mathcal{N}(z_1|\mu_{z_1}, \sigma_{z_1}^2) \mathcal{N}(z_2|\mu_{z_2}, \sigma_{z_2}^2)$





#### Detour

- Consider a scalar transformation of a scalar random variable u as  $\theta = T(u)$
- ${\ }^{\bullet}$  Probability distributions of random variables u and heta are related as

$$p(\theta) = p(u) \left| \frac{du}{d\theta} \right|$$

• Similarly, for multivariate random variables (of same size) related as  $\theta = T(u)$ 

$$p(\boldsymbol{\theta}) = p(\boldsymbol{u}) \left| \det \left( \frac{\partial \boldsymbol{u}}{\partial \boldsymbol{\theta}} \right) \right|^{Absolute value of the determinant of the Jacobian (note that \boldsymbol{u} = T^{-1}(\boldsymbol{\theta}))^{-1}$$

• We can use such transformations for VI by using a simple distribution for q(Z) and then transform it to a more expressive/appropriate distribution (more on this later)

A one-to-one

transformation function

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Transformed

random variable

## Mean-Field VI

- A special way to maximize the ELBO is via the mean-field approximation
- Doesn't require specifying the form of  $q(\mathbf{Z}|\phi)$  or computing ELBO's gradients
- The idea: Assumes unknowns Z can be partitioned into M groups  $Z_1, Z_2, \ldots, Z_M$ , s.t.,

For models with local conjugacy, it becomes super easy!

- Learning the optimal  $q(\mathbf{Z}|\phi)$  reduces to learning the optimal  $q_1, q_2, \dots, q_M$
- Can select groupsbased on model's structure, e.g., in Bayesian neural net for regression

$$p(w|X,y,\lambda,eta) pprox q(w|\phi) = \prod_{\ell=1}^L q(w^{(\ell)}|\phi_\ell) - Assuming a network with L ayers, mean-field across layers are an event of the second sec$$

- Mean-field has limitations. Factorized form ignores the correlations among unknowns
  - Variants such as "structured mean-field" exist where some correlations can be modeled

# Deriving Mean-Field VI Updates

Writing this is the same as  $\operatorname{argmax}_{\phi} \mathcal{L}(\phi, \Theta)$ . We are just writing optimization w.r.t. q directly

- With  $q = \prod_{i=1}^{M} q_i$ , what's the optimal  $q_i$  when we do  $\operatorname{argmax}_q \mathcal{L}(q)$ ?
- Note that under this mean-field assumption, the ELBO simplifies to

$$\mathcal{L}(q) = \int q(\mathbf{Z}) \log \left[ \frac{p(\mathcal{D}, \mathbf{Z}|\Theta)}{q(\mathbf{Z})} \right] d\mathbf{Z} = \int \prod_{i} q_{i} \left[ \log p(\mathcal{D}, \mathbf{Z}|\Theta) - \sum_{i} \log q_{i} \right] d\mathbf{Z}$$
• Suppose we wish to find the optimal  $q_{j}$  given all other  $q_{i}$ 's  $(i \neq j)$  as fixed, then
$$\mathcal{L}(q) = \int q_{j} \left[ \int \log p(\mathcal{D}, \mathbf{Z}|\Theta) \prod_{i\neq j} q_{i} dZ_{i} \right] dZ_{j} - \int q_{j} \log q_{j} dZ_{j} + \text{const w. r. t. } q_{j}$$

$$= \int q_{j} \log \hat{p}(\mathcal{D}, Z_{j}|\Theta) dZ_{j} - \int q_{j} \log q_{j} Z_{j}$$

$$= -\text{KL}(q_{j}||\hat{p}) \log \hat{p}(\mathcal{D}, Z_{j}|\Theta) = \mathbb{E}_{i\neq j}[\log p(\mathcal{D}, \mathbf{Z}|\Theta]] + \text{const}$$
• Thus  $q_{j}^{*} = \operatorname{argmax}_{q_{j}} \mathcal{L}(q) = \operatorname{argmin}_{q_{j}} \text{KL}(q_{j}||\hat{p}) = \hat{p}(\mathcal{D}, Z_{j}|\Theta)$ 

# Deriving Mean-Field VI Updates

• So we saw that the optimal  $q_i$  when doing mean-field VI is

$$q_j^*(\mathbf{Z}_j) = \frac{\exp(\mathbb{E}_{i\neq j}[\log p(\mathbf{\mathcal{D}}, \mathbf{Z}|\Theta)])}{\int \exp(\mathbb{E}_{i\neq j}[\log p(\mathbf{\mathcal{D}}, \mathbf{Z}|\Theta)] d\mathbf{Z}_j}$$

- Note: Can often just compute the numerator and recognize denominator by inspection
- Important: For locally conj models,  $q_j^*(Z_j)$  will have the same form as prior  $p(Z_j|\Theta)$ 
  - Only the distribution parameters will be different
- Important: For estimating  $q_j$  the required expectation depends on other  $\{q_i\}_{i\neq j}$ 
  - Thus we use an alternating update scheme for these
- Guaranteed to converge (to a local optima)
  - We are basically solving a sequence of concave maximization problems
  - Reason:  $\mathcal{L}(q) = \int q_j \log \hat{p}(\mathcal{D}, Z_j | \Theta) Z_j \int q_j \log q_j Z_j$  is concave in  $q_j$



## The Mean-Field VI Algorithm

- Also known as Co-ordinate Ascent Variational Inference (CAVI) Algorithm
- Input: Model in form of priors and likelihood, or joint  $p(\mathcal{D}, Z | \Theta)$ , Data  $\mathcal{D}$
- Output: A variational distribution  $q(\mathbf{Z}) = \prod_{j=1}^{M} q_j(\mathbf{Z}_j)$
- Initialize: Variational distributions  $q_j(\mathbf{Z}_j)$ , j = 1, 2, ..., M
- While the ELBO has not converged
  - For each j = 1, 2, ..., M, set

 $q_j(\mathbf{Z}_j) \propto \exp(\mathbb{E}_{i\neq j}[\log p(\mathcal{D}, \mathbf{Z}|\Theta)])$ 

- Compute ELBO  $\mathcal{L}(q) = \mathbb{E}_q[\log p(\mathcal{D}, \mathbf{Z}|\Theta)] \mathbb{E}_q[\log q(\mathbf{Z})]$
- NOTE: We can also use mean-field assumption for q(Z) and optimize the ELBO using gradient based methods if we don't have local conjugacy

# VI and Convergence

- VI is guaranteed to converge to a local optima (just like EM)
- Therefore proper initialization is important (just like EM)
  - Can sometimes run multiple times with different initializations and choose the best run



- ELBO increases monotonically with iterations
  - Can thus monitor the ELBO to assess convergence



## ELBO for Model Selection

- Recall that ELBO is a <u>lower bound</u> on log of model evidence  $\log p(X|m)$
- Can compute ELBO for each model m and choose the one with largest ELBO

Plot of the variational lower bound  $\mathcal{L}$  versus the number K of components in the Gaussian mixture model, for the Old Faithful data, showing a distinct peak at K =2 components. For each value of K, the model is trained from 100 different random starts, and the results shown as '+' symbols  $p(\mathcal{D}|K)$ plotted with small random horizontal perturbations so that they can be distinguished. Note that some solutions find suboptimal local maxima, but that this happens infrequently.



Some criticism since we are using a lower-bound but often works well in practice

# VI might <u>under-estimate</u> posterior's variance

• Recall that VI approximates a posterior p by finding q that minimizes KL(q||p)

$$\mathrm{KL}(q||p) = -\int q(\mathbf{Z})\log\left\{\frac{p(\mathbf{Z}|\mathcal{D})}{q(\mathbf{Z})}\right\}d\mathbf{Z}$$

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- q(Z) will be small where  $p(Z|\mathcal{D})$  is small otherwise KL will blow up
- Thus  $q(\mathbf{Z})$  avoids low-probability regions of the true posterior



## Variational EM

- If the parameters  $\Theta$  are also unknown then we can use variational EM (VEM)
- VEM is the same as EM except the E step uses VI to approximate the CP of Z
- VEM alternates between the following two steps
  - Maximize the ELBO w.r.t.  $\phi$  (gives the variational approximation q(Z) of CP of Z)

 $\phi^{(t)} = \operatorname{argmax}_{\phi} \mathbb{E}_{q_{\phi}(Z)} \left[ \log p(\mathcal{D}, Z | \Theta^{(t-1)}) - \log q_{\phi}(Z) \right]$ 

• Maximize the ELBO w.r.t.  $\Theta$  (gives us point estimate of  $\Theta$ )

 $\Theta^{(t)} = \operatorname{argmax}_{\Theta} \mathbb{E}_{q_{\phi^{(t)}}(Z)} \left[ \log p(\mathcal{D}, Z | \Theta) - \log q_{\phi^{(t)}}(Z) \right]$ 

 $= \operatorname{argmax}_{\Theta} \mathbb{E}_{q_{\phi}(t)(Z)}[\log p(\mathcal{D}, Z|\Theta)] \stackrel{\text{This looks very similar to the}}{=} \operatorname{expected CLL with the CP rep}$ 

expected CLL with the CP replaced by its variational approximation

• Note: If we want posterior for  $\Theta$  as well, treat it similar to Z and apply variational approximation (instead of using VEM) if the posterior isn't tractable CS772A: PML

#### Extra Slides - Mean-Field VI: A Simple Example

- Consider data  $\mathbf{X} = \{x_1, x_2, \dots, x_N\}$  from a one-dim Gaussian  $\mathcal{N}(\mu, \tau^{-1})$
- Assume the following normal-gamma prior on  $\mu$  and au

 $p(\mu|\tau) = \mathcal{N}(\mu|\mu_0, (\lambda_0\tau)^{-1}) \quad p(\tau) = \text{Gamma}(\tau|a_0, b_0)$ 

- Posterior is also normal-gamma due to the jointly conjugate prior
- Let's anyway verify this by trying mean-field VI for this model
- With mean-field assumption on the variational posterior  $q(\mu, \tau) = q_{\mu}(\mu)q_{\tau}(\tau)$

$$\log q_{\mu}^{*}(\mu) = \mathbb{E}_{q_{\tau}}[\log p(\mathbf{X}, \mu, \tau)] + \text{const}$$
$$\log q_{\tau}^{*}(\tau) = \mathbb{E}_{q_{\mu}}[\log p(\mathbf{X}, \mu, \tau)] + \text{const}$$

• In this example, the log-joint  $\log p(\mathbf{X}, \mu, \tau) = \log p(\mathbf{X}|\mu, \tau) + \log p(\mu|\tau) + \log p(\tau)$ . Thus

 $\log q_{\mu}^{*}(\mu) = \mathbb{E}_{q_{\mu}}[\log p(\mathbf{X}|\mu,\tau) + \log p(\mu|\tau)] + \text{const} \quad (\text{only keeping terms that involve } \mu)$  $\log q_{\tau}^{*}(\tau) = \mathbb{E}_{q_{\mu}}[\log p(\mathbf{X}|\mu,\tau) + \log p(\mu|\tau) + \log p(\tau)] + \text{const}$ 

#### Extra Slides - Mean-Field VI: A Simple Example

• Substituting  $p(\mathbf{X}|\mu,\tau) = \prod_{n=1}^{N} p(x_n|\mu,\tau)$  and  $p(\mu|\tau)$ , we get

$$\log q_{\mu}^{*}(\mu) = \mathbb{E}_{q_{\tau}}[\log p(\mathbf{X}|\mu,\tau) + \log p(\mu|\tau)] + \text{const}$$
$$= -\frac{\mathbb{E}_{q_{\tau}}[\tau]}{2} \left\{ \sum_{n=1}^{N} (x_{n}-\mu)^{2} + \lambda_{0}(\mu-\mu_{0})^{2} \right\} + \text{const}$$

• (Verify) The above is log of a Gaussian. This  $q_{\mu}^* = \mathcal{N}(\mu | \mu_N, \lambda_N)$  with

$$\mu_N = \frac{\lambda_0 \mu_0 + N \bar{x}}{\lambda_0 + N} \quad \text{and} \quad \lambda_N = (\lambda_0 + N) \mathbb{E}_{q_\tau} [\tau]^{2}$$
 This update depends on  $q_\tau$ 

• Proceeding in a similar way (verify), we can show that  $q_{\tau}^* = \text{Gamma}(\tau | a_N, b_N)$ 

$$a_N = a_0 + rac{N+1}{2}$$
 and  $b_N = b_0 + rac{1}{2} \mathbb{E}_{q_\mu} \left[ \sum_{n=1}^N (x_n - \mu)^2 + \lambda_0 (\mu - \mu_0)^2 \right]^2$  This update depends on  $q_\mu$ 

• Note: Updates of  $q_{\mu}^{*}$  and  $q_{\tau}^{*}$  depend on each other (hence alternating updates needed)