

Computing the Posterior in Probabilistic Linear Regression

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Inferring the Posterior Distribution (fully Bayesian Inference)

- Inferring the full posterior is straightforward if the hyperparams β and λ to be known/fixed
 - Basically, the conjugacy helps here (Gaussian prior is **conjugate** to Gaussian likelihood)
- The posterior over the weight vector \mathbf{w} (with β and λ known)

$$p(\mathbf{w}|\mathbf{X}, \mathbf{y}, \beta, \lambda) = \frac{p(\mathbf{y}|\mathbf{X}, \mathbf{w}, \beta)p(\mathbf{w}|\lambda)}{p(\mathbf{y}|\mathbf{X}, \beta, \lambda)}$$

- Computing $P(\mathbf{w}|\mathbf{X}, \mathbf{y}, \beta, \lambda)$ (like Bernoulli-Beta case, doing it only upto proportionality constant)

$$P(\mathbf{w}|\mathbf{X}, \mathbf{y}, \beta, \lambda) \propto P(\mathbf{w}|\lambda)P(\mathbf{y}|\mathbf{X}, \mathbf{w}, \beta)$$

- After some algebra, this gets simplified into the following (proof on the next two slides)

$$P(\mathbf{w}|\mathbf{X}, \mathbf{y}, \beta, \lambda) = \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \quad (\text{The posterior must be Gaussian due to conjugacy})$$

$$\text{where } \boldsymbol{\Sigma} = \left(\beta \sum_{n=1}^N \mathbf{x}_n \mathbf{x}_n^T + \lambda \mathbf{I}_D \right)^{-1} = \left(\beta \mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}_D \right)^{-1}$$

$$\boldsymbol{\mu} = \boldsymbol{\Sigma} \left(\beta \sum_{n=1}^N y_n \mathbf{x}_n \right) = \boldsymbol{\Sigma} (\beta \mathbf{X}^T \mathbf{y}) = \left(\mathbf{X}^T \mathbf{X} + \frac{\lambda}{\beta} \mathbf{I}_D \right)^{-1} \mathbf{X}^T \mathbf{y}$$



The “Completing The Square” Trick for Gaussian Posterior

- Plugging in the respective distributions for $p(\mathbf{w}|\lambda)$ and $p(\mathbf{y}|\mathbf{X}, \mathbf{w}, \beta)$, we will get

$$\begin{aligned} p(\mathbf{w}|\mathbf{X}, \mathbf{y}, \beta, \lambda) &\propto p(\mathbf{w}|\lambda)p(\mathbf{y}|\mathbf{X}, \mathbf{w}, \beta) = \mathcal{N}(\mathbf{w}|\mathbf{0}, \lambda^{-1}\mathbf{I}_D)\mathcal{N}(\mathbf{y}|\mathbf{X}\mathbf{w}, \beta^{-1}\mathbf{I}_N) \\ &\propto \exp\left(-\frac{\lambda}{2}\mathbf{w}^\top\mathbf{w}\right)\exp\left(-\frac{\beta}{2}(\mathbf{y}-\mathbf{X}\mathbf{w})^\top(\mathbf{y}-\mathbf{X}\mathbf{w})\right) \\ &= \exp\left[-\frac{\lambda}{2}\mathbf{w}^\top\mathbf{w} - \frac{\beta}{2}(\mathbf{y}^\top\mathbf{y} + \mathbf{w}^\top\mathbf{X}^\top\mathbf{X}\mathbf{w} - 2\mathbf{w}^\top\mathbf{X}^\top\mathbf{y})\right] \\ &\propto \exp\left[-\frac{\lambda}{2}\mathbf{w}^\top\mathbf{w} - \frac{\beta}{2}(\mathbf{w}^\top\mathbf{X}^\top\mathbf{X}\mathbf{w} - 2\mathbf{w}^\top\mathbf{X}^\top\mathbf{y})\right] \\ &= \exp\left[-\frac{1}{2}\left(\mathbf{w}^\top(\lambda\mathbf{I}_D + \beta\mathbf{X}^\top\mathbf{X})\mathbf{w} - 2\beta\mathbf{w}^\top\mathbf{X}^\top\mathbf{y}\right)\right] \end{aligned}$$

- We will now try to bring the exponent into a quadratic form to see if it corresponds to some Gaussian. So basically, we will use the “complete the square” trick



The “Completing The Square” Trick for Gaussian Posterior

- So we had.. $p(\mathbf{w}|\mathbf{X}, \mathbf{y}, \beta, \lambda) \propto \exp \left[-\frac{1}{2} \left(\mathbf{w}^\top (\lambda \mathbf{I}_D + \beta \mathbf{X}^\top \mathbf{X}) \mathbf{w} - 2\beta \mathbf{w}^\top \mathbf{X}^\top \mathbf{y} \right) \right]$
- Let's see if we can bring the above posterior into the form of the following Gaussian

$$\mathcal{N}(\mathbf{w}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) \propto \exp \left[-\frac{1}{2} (\mathbf{w} - \boldsymbol{\mu})^\top \boldsymbol{\Sigma}^{-1} (\mathbf{w} - \boldsymbol{\mu}) \right] = \exp \left[-\frac{1}{2} (\mathbf{w}^\top \boldsymbol{\Sigma}^{-1} \mathbf{w} - 2\mathbf{w}^\top \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu} + \boldsymbol{\mu}^\top \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}) \right]$$

- Let's multiply and divide $p(\mathbf{w}|\mathbf{X}, \mathbf{y}, \beta, \lambda) \propto \exp \left[-\frac{1}{2} \left(\mathbf{w}^\top (\lambda \mathbf{I}_D + \beta \mathbf{X}^\top \mathbf{X}) \mathbf{w} - 2\beta \mathbf{w}^\top \mathbf{X}^\top \mathbf{y} \right) \right]$ by $\exp \left[-\frac{1}{2} \boldsymbol{\mu}^\top \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu} \right]$
- This gives the following up to a prop. constant (remember $\boldsymbol{\mu}^\top \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}$ is constant w.r.t. \mathbf{w}):

$$p(\mathbf{w}|\mathbf{X}, \mathbf{y}, \beta, \lambda) \propto \exp \left[-\frac{1}{2} \left(\mathbf{w}^\top (\lambda \mathbf{I}_D + \beta \mathbf{X}^\top \mathbf{X}) \mathbf{w} - 2\beta \mathbf{w}^\top \mathbf{X}^\top \mathbf{y} + \boldsymbol{\mu}^\top \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu} \right) \right]$$

- Finally comparing with the expression of $\mathcal{N}(\mathbf{w}|\boldsymbol{\mu}, \boldsymbol{\Sigma})$ we can see that

$$\begin{aligned} \boldsymbol{\Sigma} &= (\lambda \mathbf{I}_D + \beta \mathbf{X}^\top \mathbf{X})^{-1} \\ \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu} &= \beta \mathbf{X}^\top \mathbf{y} \Rightarrow \boldsymbol{\mu} = \boldsymbol{\Sigma} (\beta \mathbf{X}^\top \mathbf{y}) = (\mathbf{X}^\top \mathbf{X} + \frac{\lambda}{\beta} \mathbf{I}_D)^{-1} \mathbf{X}^\top \mathbf{y} \end{aligned}$$

- Note: The above expression for the posterior can also be directly obtained using properties of Gaussian distributions (Refer to the maths refresher slides on “reverse conditionals”, or MLAPP 4.3-4.4)

