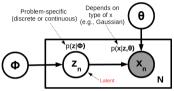
### Latent Variable Models and Expectation Maximization

Piyush Rai

#### Introduction to Machine Learning (CS771A)

September 27, 2018

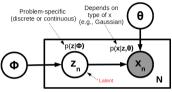
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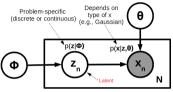


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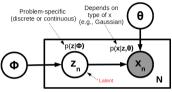
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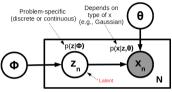
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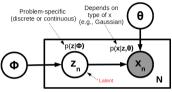
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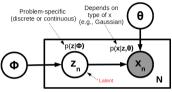
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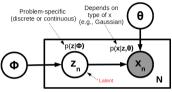
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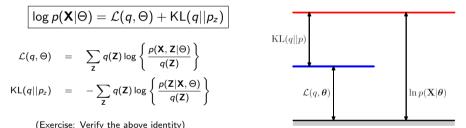
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$$\begin{bmatrix} \log p(\mathbf{X}|\Theta) = \mathcal{L}(q,\Theta) + \mathsf{KL}(q||p_z) \\ \mathcal{L}(q,\Theta) = \sum_{\mathbf{Z}} q(\mathbf{Z}) \log \left\{ \frac{p(\mathbf{X},\mathbf{Z}|\Theta)}{q(\mathbf{Z})} \right\} \\ \mathsf{KL}(q||p_z) = -\sum_{\mathbf{Z}} q(\mathbf{Z}) \log \left\{ \frac{p(\mathbf{Z}|\mathbf{X},\Theta)}{q(\mathbf{Z})} \right\} \\ \text{(Exercise: Verify the above identity)} \end{bmatrix}$$

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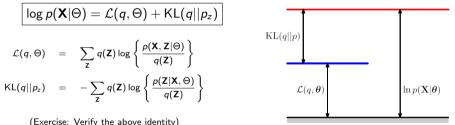
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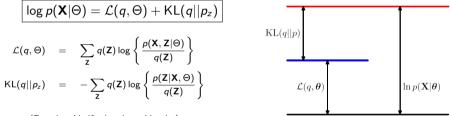
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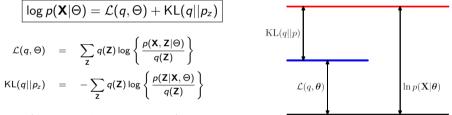
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• Since  $\mathsf{KL}(q||p_z) \geq$  0,  $\mathcal{L}(q,\Theta)$  is a lower-bound on log  $p(\mathbf{X}|\Theta)$ 

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• Maximizing  $\mathcal{L}(q, \Theta)$  will also improve  $\log p(\mathbf{X}|\Theta)$ . Also, as we'll see, it's easier to maximize  $\mathcal{L}(q, \Theta)$ 

• Note that  $\mathcal{L}(q, \Theta)$  depends on two things  $q(\mathbf{Z})$  and  $\Theta$ . Let's do ALT-OPT for these



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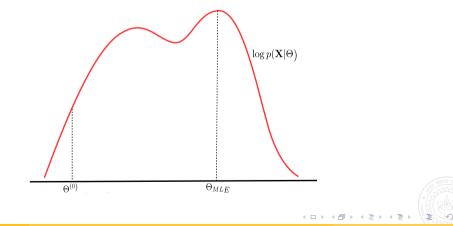
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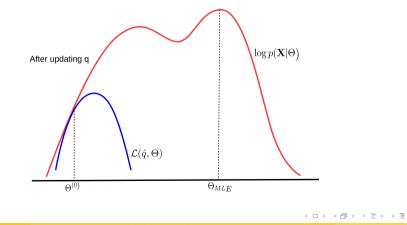
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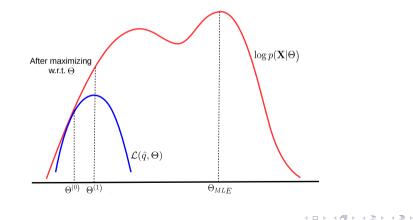
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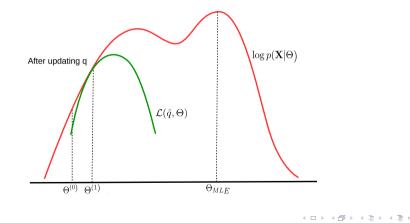
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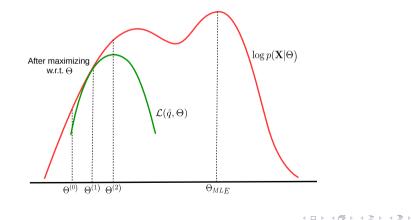
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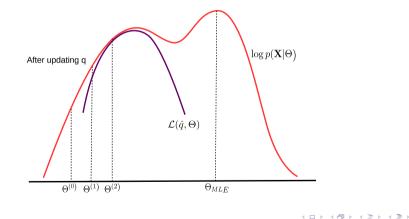
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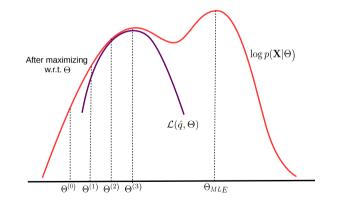
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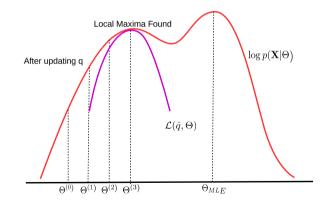
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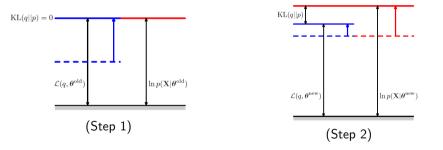


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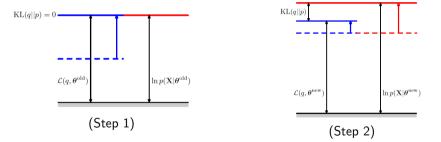
### What's Going On: Another Illustration

- The two-step alternating optimization scheme we saw can never decrease  $p(\mathbf{X}|\Theta)$  (good thing)
- To see this consider both steps: (1) Optimize q given  $\Theta = \Theta^{old}$ ; (2) Optimize  $\Theta$  given this q



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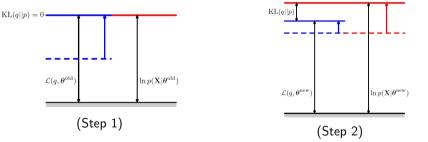
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- Step 1 keeps  $\Theta$  fixed, so  $p(\mathbf{X}|\Theta)$  obviously can't decrease (stays unchanged in this step)
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The ALT-OPT of  $\mathcal{L}(q,\Theta)$  that we saw leads to the EM algorithm (Dempster, Laird, Rubin, 1977)

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$$p(\boldsymbol{z}_n^{(t)}|\boldsymbol{x}_n, \Theta^{(t-1)}) = \frac{p(\boldsymbol{z}_n^{(t)}|\Theta^{(t-1)})p(\boldsymbol{x}_n|\boldsymbol{z}_n^{(t)}, \Theta^{(t-1)})}{p(\boldsymbol{x}_n|\Theta^{(t-1)})} \propto \text{prior} \times \text{likelihood}$$

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4

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#### The EM Algorithm

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Note: If we can take the MAP estimate  $\hat{z}_n$  of  $z_n$  (not full posterior) in Step 1 and maximize the CLL in Step 2 using that estimate, i.e., do arg max $_{\Theta} \sum_{n=1}^{N} \log p(x_n, \hat{z}_n^{(t)} | \Theta)$ , this will be identical to ALT-OPT

• Deriving the EM algorithm for any model requires finding the expression of the expected CLL

$$\begin{aligned} \mathcal{Q}(\Theta, \Theta^{old}) &= \sum_{n=1}^{N} \mathbb{E}_{p(\boldsymbol{z}_n | \boldsymbol{x}_n, \Theta^{old})} [\log p(\boldsymbol{x}_n, \boldsymbol{z}_n | \Theta)] \\ &= \sum_{n=1}^{N} \mathbb{E}_{p(\boldsymbol{z}_n | \boldsymbol{x}_n, \Theta^{old})} [\log p(\boldsymbol{x}_n | \boldsymbol{z}_n, \Theta) + \log p(\boldsymbol{z}_n | \Theta)] \end{aligned}$$



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- If  $p(\mathbf{x}_n | \mathbf{z}_n, \Theta)$  and  $p(\mathbf{z}_n | \Theta)$  are exp-family distributions, expected CLL will have a simple form
- Finding the expression for the expected CLL in such cases is fairly straightforward
  - First write down the expressions for  $p(x_n|z_n,\Theta)$  and  $p(z_n|\Theta)$  and simplify as much as possible

• Deriving the EM algorithm for any model requires finding the expression of the expected CLL

$$\begin{aligned} \mathcal{Q}(\Theta, \Theta^{old}) &= \sum_{n=1}^{N} \mathbb{E}_{p(\boldsymbol{z}_{n} | \boldsymbol{x}_{n}, \Theta^{old})}[\log p(\boldsymbol{x}_{n}, \boldsymbol{z}_{n} | \Theta)] \\ &= \sum_{n=1}^{N} \mathbb{E}_{p(\boldsymbol{z}_{n} | \boldsymbol{x}_{n}, \Theta^{old})}[\log p(\boldsymbol{x}_{n} | \boldsymbol{z}_{n}, \Theta) + \log p(\boldsymbol{z}_{n} | \Theta)] \end{aligned}$$

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    - .. and so on..
- The expected CLL may not always be computable and may need to be approximated



• Let's first look at the CLL. Similar to generative classification with Gaussian class-conditionals

$$\log p(\mathbf{X}, \mathbf{Z}|\Theta) = \sum_{n=1}^{N} \sum_{k=1}^{K} z_{nk} [\log \pi_k + \log \mathcal{N}(\mathbf{x}_n | \mu_k, \Sigma_k)] \qquad \text{(we've seen how we get this)}$$



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.. where the expectation is w.r.t. the current posterior of  $z_n$ , i.e.,  $p(z_n|x_n, \Theta^{old})$ 

• In this case, we only need  $\mathbb{E}[z_{nk}]$  which can be computed as

$$\mathbb{E}[z_{nk}] = \gamma_{nk} \quad = \quad \mathbf{0} \times p(z_{nk} = \mathbf{0} | \mathbf{x}_n, \Theta^{old}) + 1 \times p(z_{nk} = 1 | \mathbf{x}_n, \Theta^{old}) = p(z_{nk} = 1 | \mathbf{x}_n)$$

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$$\propto p(z_{nk} = 1)p(\mathbf{x}_n | z_{nk} = 1)$$
 (from Bayes Rule)

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• Note: We can finally normalize  $\mathbb{E}[z_{nk}]$  as  $\mathbb{E}[z_{nk}] = \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{\ell=1}^K \pi_\ell \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_\ell, \boldsymbol{\Sigma}_\ell)}$  since  $\sum_{k=1}^K \mathbb{E}[z_{nk}] = 1$ 

EM for Gaussian Mixture Model



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• Initialize  $\Theta = \{\pi_k, \mu_k, \Sigma_k\}_{k=1}^K$  as  $\Theta^{(0)}$ , set t = 1

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$$\mathbb{E}[\boldsymbol{z}_{nk}^{(t)}] = \gamma_{nk}^{(t)} = \frac{\pi_k^{(t-1)} \mathcal{N}(\boldsymbol{x}_n | \boldsymbol{\mu}_k^{(t-1)}, \boldsymbol{\Sigma}_k^{(t-1)})}{\sum_{\ell=1}^{K} \pi_\ell^{(t-1)} \mathcal{N}(\boldsymbol{x}_n | \boldsymbol{\mu}_\ell^{(t-1)}, \boldsymbol{\Sigma}_\ell^{(t-1)})} \quad \forall n, k \in \mathbb{C}$$

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• Given "responsibilities"  $\gamma_{nk} = \mathbb{E}[z_{nk}]$ , and  $N_k = \sum_{n=1}^N \gamma_{nk}$ , re-estimate  $\Theta$  via MLE

$$\boldsymbol{\mu}_{k}^{(t)} = \frac{1}{N_{k}} \sum_{n=1}^{N} \boldsymbol{\gamma}_{nk}^{(t)} \boldsymbol{x}_{nk}$$

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$$\begin{split} \boldsymbol{\mu}_{k}^{(t)} &= \frac{1}{N_{k}} \sum_{n=1}^{N} \gamma_{nk}^{(t)} \boldsymbol{x}_{n} \\ \boldsymbol{\Sigma}_{k}^{(t)} &= \frac{1}{N_{k}} \sum_{n=1}^{N} \gamma_{nk}^{(t)} (\boldsymbol{x}_{n} - \boldsymbol{\mu}_{k}^{(t)}) (\boldsymbol{x}_{n} - \boldsymbol{\mu}_{k}^{(t)})^{\mathsf{T}} \end{split}$$

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$$\mu_{k}^{(t)} = \frac{1}{N_{k}} \sum_{n=1}^{N} \gamma_{nk}^{(t)} \mathbf{x}_{n}$$

$$\mathbf{\Sigma}_{k}^{(t)} = \frac{1}{N_{k}} \sum_{n=1}^{N} \gamma_{nk}^{(t)} (\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(t)}) (\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(t)})^{-1}$$

$$\pi_{k}^{(t)} = \frac{N_{k}}{N_{k}}$$

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$$\boldsymbol{\mu}_{k}^{(t)} = \frac{1}{N_{k}} \sum_{n=1}^{N} \gamma_{nk}^{(t)} \mathbf{x}_{n}$$
$$\boldsymbol{\Sigma}_{k}^{(t)} = \frac{1}{N_{k}} \sum_{n=1}^{N} \gamma_{nk}^{(t)} (\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(t)}) (\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(t)})^{\top}$$
$$\boldsymbol{\pi}_{k}^{(t)} = \frac{N_{k}}{2}$$

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• Set t = t + 1 and go to step 2 if not yet converged

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• Let's consider a latent factor model for dimensionality reduction (will revisit this later)

$$p(\mathbf{x}_n | \mathbf{z}_n, \mathbf{W}, \sigma^2) = \mathcal{N}(\mathbf{W}\mathbf{z}_n, \sigma^2 \mathbf{I}_D) \qquad p(\mathbf{z}_n) = \mathcal{N}(\mathbf{0}, \mathbf{I}_K)$$



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• A low-dim  $\boldsymbol{z}_n \in \mathbb{R}^K$  mapped to high-dim  $\boldsymbol{x}_n \in \mathbb{R}^D$  via a projection matrix  $\boldsymbol{\mathsf{W}} \in \mathbb{R}^{D imes K}$ 



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- The complete data log-likelihood for this model will be

 $\log p(\mathbf{X}, \mathbf{Z} | \mathbf{W}, \sigma^2)$ 

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$$\log p(\mathbf{X}, \mathbf{Z} | \mathbf{W}, \sigma^2) = \log \prod_{n=1}^{N} p(\mathbf{x}_n, \mathbf{z}_n | \mathbf{W}, \sigma^2)$$

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$$p(\mathbf{x}_n | \mathbf{z}_n, \mathbf{W}, \sigma^2) = \mathcal{N}(\mathbf{W}\mathbf{z}_n, \sigma^2 \mathbf{I}_D) \qquad p(\mathbf{z}_n) = \mathcal{N}(\mathbf{0}, \mathbf{I}_K)$$

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$$\log p(\mathbf{X}, \mathbf{Z} | \mathbf{W}, \sigma^2) = \log \prod_{n=1}^{N} p(\mathbf{x}_n, \mathbf{z}_n | \mathbf{W}, \sigma^2) = \log \prod_{n=1}^{N} p(\mathbf{x}_n | \mathbf{z}_n, \mathbf{W}, \sigma^2) p(\mathbf{z}_n)$$

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• Plugging in the expressions for  $p(\mathbf{x}_n | \mathbf{z}_n, \mathbf{W}, \sigma^2)$  and  $p(\mathbf{z}_n)$  and simplifying (exercise)

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• Let's consider a latent factor model for dimensionality reduction (will revisit this later)

$$p(\mathbf{x}_n | \mathbf{z}_n, \mathbf{W}, \sigma^2) = \mathcal{N}(\mathbf{W}\mathbf{z}_n, \sigma^2 \mathbf{I}_D) \qquad p(\mathbf{z}_n) = \mathcal{N}(\mathbf{0}, \mathbf{I}_K)$$

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- The M step maximizes the expected CLL w.r.t. the parameters ( $\mathbf{W}, \sigma^2$  in this case)

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